1 Meta-analysis of fecal metagenomes reveals global microbial signatures that

2 are specific for colorectal cancer

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67 ABSTRACT

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69 Association studies have linked microbiome alterations with many human diseases, but not always 70 reported consistent results, which necessitates cross-study comparisons. Here, a meta-analysis of 71 eight geographically and technically diverse fecal shotgun metagenomic studies of colorectal cancer 72 (CRC, N = 768), which was controlled for several confounders, identified a core set of 29 species 73 significantly enriched in CRC metagenomes (FDR < 1E-5). CRC signatures derived from single 74 studies maintained accuracy in other studies. By training on multiple studies we improved detection 75 accuracy and disease specificity for CRC. Functional analysis of CRC metagenomes revealed 76 enriched protein and mucin catabolism genes and depleted carbohydrate degradation genes. 77 Moreover we inferred elevated production of secondary bile acids from CRC metagenomes 78 suggesting a metabolic link between cancer-associated gut microbes and a fat- and meat-rich diet. 79 Through extensive validations, this meta-analysis firmly establishes globally generalizable, predictive 80 taxonomic and functional microbiome CRC signatures as a basis for future diagnostics.

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83 INTRODUCTION

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85 Studying microbial communities colonizing the human body in a culture-independent manner has been 86 enabled by metagenomic sequencing technologies [1]. These have yielded glimpses into the complex 87 vet incompletely understood interactions between the gut microbiome - the microbial ecosystem 88 residing primarily in the large intestine - and its host [2]. To explore microbiome-host interactions in a 89 disease context, metagenome-wide association studies (MWAS) have begun to map gut microbiome 90 alterations in diabetes, inflammatory bowel disease, colorectal cancer and many other conditions [3-91 12]. However, due to the many biological factors possibly influencing gut microbiome composition in 92 addition to the condition studied, a current challenge for MWAS is confounding, which can cause false 93 associations [13, 14]. This issue is further aggravated by a lack of standards in metagenomic data 94 generation and processing, making it difficult to disentangle technical from biological effects [15].

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96 Robustness of microbiome-disease associations can be assessed through comparisons across 97 multiple metagenomic case-control studies, i.e. meta-analyses. These aim at identifying associations 98 that are consistent across studies and thus less likely attributable to biological or technical 99 confounders. Most informative are meta-analyses of populations from diverse geographic and cultural 100 regions. Previous microbiome meta-analyses based on 16S rRNA gene amplicon data found stark 101 technical differences between studies and the reported taxonomic disease associations were either of 102 low effect size or not well resolved [16-18]. In contrast, shotgun metagenomics enables analyses with 103 higher taxonomic resolution and of gene functions to improve statistical power for fine-mapping 104 disease-associated strains and aid in the interpretation of host-microbial co-metabolism. Thus far 105 however, meta-analyses of shotgun metagenomic data have either reported on features of general 106 dysbiosis in comparisons across multiple diseases [19], or have left it unclear how well microbiome

signatures generalize across studies of the same disease when data are rigorously separated to avoid
 over-optimistic evaluations of their prediction accuracy [20].

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110 Here, we present a meta-analysis of a total of eight studies of CRC including fecal metagenomic data 111 from 386 cancer cases and 392 tumor-free controls. After consistent data reprocessing, we examined 112 an initial set of five studies for CRC-associated changes in the gut microbiome. Firstly, we investigated 113 potential confounders, followed by identifying (univariate) microbial species associations, and inferring 114 species co-occurrence patterns in CRC. Secondly, we trained multivariable classification models for 115 recognition of CRC status, from both taxonomic and functional microbiome profiles and tested how 116 accurately these models generalized to data from studies not used for training. Moreover, we 117 evaluated performance improvements achieved by pooling data across studies and the disease-118 specificity of the resulting classification models. Thirdly, targeted investigation of virulence and toxicity 119 genes as candidate functional biomarkers for CRC revealed several of these to be enriched in CRC 120 metagenomes indicative of their prevalence and potential relevance in CRC patients. Three additional, 121 more recent studies were finally used to independently validate these taxonomic and functional CRC 122 signatures.

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124 RESULTS

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126 Consistent processing of published and new data for meta-analysis of CRC metagenomes

127 In this meta-analysis we included four published studies which used fecal shotgun metagenomics to 128 characterize CRC patients compared to healthy controls (referred to by the country codes FR, AT, CN, 129 and US, corresponding to the respective main study population; see Table 1, Supplementary Table 130 S1, and Methods for inclusion criteria). For an additional fifth study population, we generated new 131 fecal metagenomic data from samples collected in Germany (herein abbreviated as DE); a subset of 132 samples from this patient collective were published previously (Table 1, Methods, [8]). These five 133 studies were conducted on three continents and differed in sampling procedures, sample storage, and 134 DNA extraction protocols. Notably, the fecal specimen of the US study were freeze-dried and stored at 135 -80°C for more than 25 years before DNA extraction and sequencing [10]. In all studies, however, 136 samples were collected prior to treatment, thus excluding cancer therapy as a potential confounding 137 effect [14, 21]. Most samples were even taken before bowel preparation for colonoscopy, with some 138 exceptions in the DE, CN and US studies (Supplementary Table S2). To ensure consistency in 139 bioinformatic analyses, all raw sequencing data were (re-)processed using mOTUs2 for taxonomic 140 profiling [22] and MOCAT2 for functional profiling [23].

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142 Univariate meta-analysis of species associated with CRC

The first aim of the meta-analysis was to determine gut microbial species that are enriched or depleted in CRC metagenomes in a consistent manner across the five study populations. However, as these studies differed from one another in many biological and technical aspects, we first quantified the effect of study-associated heterogeneity on microbiome composition. We contrasted this with other potential confounders ('patient age', 'BMI', 'sex', 'sampling after colonoscopy', and 'library size'; 148 additionally, 'smoking status', 'type II diabetes comorbidity', and 'vegetarian diet' where available 149 Extended Data 1, Supplementary Table S3). This analysis revealed the factor 'study' to have a 150 predominant impact on species composition, which is supported by a recent comparison of DNA 151 extraction protocols, as these typically differ between studies [15]. An analysis of microbial alpha and 152 beta diversity showed study heterogeneity to also have a larger effect on overall microbiome 153 composition than CRC in our data (Extended Data 2).

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155 For the identification of microbial taxa significantly differing in abundance in CRC, parametric effect 156 size measures are not well established, because microbiome data is characterized by non-Gaussian 157 distributions with extreme dispersion; we thus used a generalisation of the fold change (Extended 158 Data 3) and non-parametric significance testing. In this permutation test framework [24] (herein 159 referred to as blocked univariate Wilcoxon tests) differential abundance in CRC can be assessed 160 while accounting for 'study' as a nuisance effect that is treated as a blocking factor; additionally, 161 motivated by our confounder analysis, we also blocked for 'colonoscopy' in all analyses (Methods, 162 Extended Data 1). To rule out spurious associations due to the compositional nature of microbial 163 relative abundance data, we additionally compared the results of this test with a method [25] 164 employing log-ratio transformation (and found highly correlated results, Supplementary Fig. 1, 165 Supplementary Table S4).

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167 At a meta-analysis false discovery rate (FDR) of 0.005, we identified 94 microbial species to be 168 differentially abundant in the CRC microbiome, out of 849 species consistently detected across 169 studies (Supplementary Table S4, Methods). Among these, we focused on a core set of the 29 most 170 significant markers (FDR < 1E-5, Fig. 1a) for further analysis. The latter included members of several 171 genera previously associated with CRC, such as Fusobacterium, Porphyromonas, Parvimonas, 172 Peptostreptococcus, Gemella, Prevotella, and Solobacterium (Fig. 1b, [8-11]), and 8 additional 173 species without genomic reference sequences (meta-mOTUs, Methods, [22]) mostly from the 174 Porphyromonas and Dialister genera and the Clostridiales order (see Extended Data 4 and 175 Supplementary Table S4 for genus-level associations). Collectively, these 29 core CRC-associated 176 species show a previously underappreciated diversity of 11 Clostridiales species to be enriched in 177 CRC (Fig. 1b). In contrast to the majority of species that are more strongly affected by study 178 heterogeneity than by CRC status, 26 out of the 29 CRC-associated species varied more by disease 179 status (Fig. 1d).

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181 All of the core CRC-associated species were enriched in patients and were often undetectable in 182 metagenomes from non-neoplastic controls. While previous studies were contradictory in the reported 183 proportion of positive versus negative associations [8, 9, 17, 20], our meta-analysis results are more 184 easily reconciled with a model in which - potentially many - gut microbes contribute to or benefit from 185 tumorigenesis than with the opposing model in which a lack of protective microbes contributes to CRC 186 development (Fig. 1b). Although these core taxonomic CRC associations were highly significant and 187 consistent, individual studies showed marked discrepancies in the species identified as significant 188 (Fig. 1a). Retrospective examination of the precision and sensitivity with which individual studies

detected this core of CRC-associated species showed relatively low sensitivity for the US study
 (consistent with the original report [10]) and low precision of the AT study due to associations that
 were not replicated in other studies (Supplementary Fig. 2).

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193 Analyzing patient metagenomes for co-occurrences among the core set of 29 species that are strongly 194 enriched in the CRC microbiome revealed four species clusters with distinct taxonomic composition 195 (Fig. 2a, Extended Data 5, Methods). Two of them showed strong taxonomic consistency: Cluster 1 196 exclusively comprised Porphyromonas spp., and cluster 4 only contained members of the Clostridiales 197 order. In contrast, the other two clusters were taxonomically more heterogeneous with cluster 3 198 grouping together the species with highest prevalence in CRC cases (all among the ten most highly 199 significant markers), consistent with a co-occurrence analysis of one of the data sets included here 200 [11]. Cluster 2 contained species with intermediate prevalence.

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202 Investigating whether these four clusters were associated with different tumor characteristics, we 203 found the Porphyromonas cluster 1 to be significantly enriched in rectal tumors (Fig. 2b), consistent 204 with the presence of superoxide dismutase genes in Porphyromonas genomes possibly conferring 205 tolerance to a more aerobic milieu in the rectum (Extended Data 5). The Clostridiales cluster 4 was 206 significantly more prevalent in female CRC patients. All species clusters showed a slight tendency 207 towards late-stage CRC (i.e. AJCC stages III and IV), but this was only significant for cluster 3. 208 Associations with patient age and BMI were weaker and not significant (Extended Data 5). To rule out 209 secondary effects due to differences in patient composition among studies, all of these tests were 210 corrected for study effects (by blocking for 'study' and 'colonoscopy', see Methods). At the level of 211 individual species, significant stage-specific enrichments could not be detected suggesting CRC-212 associated microbiome changes to be less dynamic during cancer progression than previously 213 postulated [26], although fecal material may be less suitable to address this question than tissue 214 samples.

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216 Metagenomic CRC classification models

To establish metagenomic signatures for CRC detection across studies in face of geographic and technical heterogeneity, we developed multivariable statistical modeling workflows with rigorous external validation to avoid prevailing issues of overfitting and over-optimistic reports of model accuracy [19]. As a precaution against over-optimistic evaluation, these workflows are independent of the above-described differential abundance analysis. Instead, LASSO (Least Absolute Shrinkage and Selection Operator) logistic regression classifiers were employed to select predictive microbial features and eliminated uninformative ones (Methods).

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In a first step, we used abundance profiles from five studies including the 849 most abundant microbial species and assessed how well classifiers trained in cross validation (CV) on one study generalize in evaluations on the other four studies (study-to-study transfer of classifiers) (**Fig. 3a**). Within-study cross-validation performance, as quantified by the Area Under the Receiver Operating Characteristics (AUROC) curve, ranged between 0.69 and 0.92 and was generally maintained in study-to-study

230 transfer (AUROC dropping by 0.07±0.12 on average) with two notable exceptions. First, in line with 231 the univariate analysis of species associations, CRC detection accuracy on the US study was lower 232 than for the other studies, both in cross-validation and in study-to-study transfer. This could potentially 233 be explained by the US fecal specimen, unlike in the other studies, being freeze-archived for >25 234 years before metagenomic sequencing [10]. Second, classifiers trained on the AT study did not 235 generalize as well to the other studies, consistent with low study precision seen in univariate meta-236 analysis (Supplementary Fig. 2). Given the microbial co-occurrence clusters described above, we 237 wondered whether species-species interactions would provide additional information relevant for CRC 238 recognition that is not contained in species abundance profiles. However, nonlinear classifiers able to 239 exploit such interactions did not yield significantly better accuracies (Supplementary Fig. 3, see also 240 [27]), suggesting that the linear model based on few biomarkers (on average 17 species account for 241 more than 80% of the classifier weight, **Extended Data 6**) is near optimal for CRC prediction.

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243 We further assessed if including data from all but one study in model training improves prediction on 244 the remaining held-out study (leave-one-study-out validation, LOSO). LOSO performance of species-245 level models ranged between 0.71 and 0.91, and when disregarding the US study as an outlier was 246 ≥0.83 (Fig. 3b). This corresponds to a LOSO accuracy increase of 0.076±0.03 compared to study-to-247 study transfer. These results suggest that one can expect a CRC detection accuracy ≥ 0.8 (AUROC) 248 for any new CRC study using similarly generated metagenomic data. We moreover verified that 249 metagenomic CRC classification models trained on species composition were not biased for clinical 250 subgroups. With the exception of slightly more sensitive detection of late stage CRC (P = 0.03, mostly 251 originating from the US study, Extended Data 7), we did not observe any classification bias by patient 252 age, sex, BMI, or localization. Together this suggests that these metagenomic classifiers are unlikely 253 to be strongly confounded by the clinical parameters recorded.

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255 Several previous studies comparing microbiome changes across multiple diseases reported primarily 256 general dysbiotic alterations and highlighted the need to examine the disease specificity of 257 microbiome signatures [17, 19]. Therefore, we assessed false positive (FP) predictions of our 258 metagenomic CRC classifiers on fecal metagenomes of type 2 diabetes [4, 5], Parkinson's disease 259 [12], ulcerative colitis and Crohn's disease [6, 7] patients, reasoning that classifiers relying on 260 biomarkers for general dysbiosis would yield an excess of FPs on these cohorts. However, our LOSO 261 classification models calibrated to have a false-positive rate (FPR) of 0.1 on CRC datasets in fact 262 maintained similarly low FPRs on other disease datasets ranging from 0.09 to 0.13 (Fig. 3c). 263 Interestingly, disease specificity of LOSO models was significantly improved over that observed for 264 classifiers trained on a single study, indicating that inclusion of multiple studies in the training set of a 265 classifier can substantially improve its specificity for a given disease.

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268 Functional metagenomic signatures for CRC

As shotgun metagenomics data, in contrast to 16S rRNA gene amplicon data, allow for a direct analysis of the functional potential of the gut microbiome, we examined how predictive metabolic 271 pathways and orthologous gene families differing in abundance between CRC patients and controls 272 would be of CRC status. When applying the same classification workflow as above to eggNOG 273 orthologous gene family abundances [28], CRC detection accuracy was very similar to that observed 274 for taxonomic models (Fig. 3de). AUROC values ranged from 0.70 to 0.81 for study-to-study transfer 275 (per-study averages, Fig. 3e) and from 0.78 to 0.89 in LOSO validation with a pattern of generalization 276 across studies resembling that for taxonomic classifiers. The accuracy of functional signatures did not 277 strongly depend on eggNOG as an annotation source, but was similar when based on other 278 comprehensive functional databases, such as KEGG [29] (Extended Data 8). When using individual 279 gene abundances from metagenomic gene catalogues as a classifier input [30], we observed higher 280 within-study cross-validation AUROC values of ≥0.96 in all studies, but lower generalization to other 281 studies (AUROC between 0.60 and 0.79) (Extended Data 8).

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283 To explore changes in metabolic capacity of gut microbiomes from CRC patients more broadly, we 284 quantified gut metabolic modules (defined in [31]) and subjected these to the same differential 285 abundance analysis developed for microbial species. Gut metabolic modules with significantly higher 286 abundance (FDR < 0.01, Wilcoxon test blocked for study and colonoscopy) in CRC metagenomes 287 predominantly belonged to pathways for the degradation of amino acids, mucins (glycoproteins) and 288 organic acids. This clear trend was accompanied by a depletion of genes from carbohydrate 289 degradation modules (Fig. 4ab). Differences in all four high-level categories were highly significant (P 290 < 1E-6 in all cases, blocked Wilcoxon tests) and consistent across studies (Fig. 4b). Overall these 291 results establish a clear shift from dietary carbohydrate utilization in a healthy gut microbiome to amino 292 acid degradation in CRC consistent with an earlier report based on a subset of the data [8]. 293 Correlation analysis suggests that increased capacity for amino acid degradation is mostly contributed 294 by CRC-associated Clostridiales (cf. cluster 4 in Fig. 2, Supplementary Fig. 4). About one half of 295 these metagenomic pathway enrichments are also in agreement with independent metabolomics data 296 suggesting increased availability of amino acids in epithelial cells or feces of CRC patients 297 (Supplementary Table S5, [32-36]). While the observed pathway enrichments could potentially result 298 from many factors, including unmeasured ones [13], they are consistent with established dietary risk 299 factors for CRC, which include red and processed meat consumption [37] and low fiber intake [38].

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301 The large metagenomic data set analyzed here allowed us to quantify the prevalence of gut microbial 302 virulence and toxicity mechanisms thought to play a role in colorectal carcinogenesis. Prominent 303 examples include the Fusobacterium nucleatum adhesion protein A (encoded by the fadA gene), the 304 Bacteroides fragilis enterotoxin (bft gene) and colibactin produced by some Escherichia coli strains 305 (pks genomic island) [39, 40]. Moreover, intestinal Clostridium spp. are known to contribute to the 306 conversion of primary to secondary bile acids using several metabolic pathways including 7a-307 dehydroxylation, encoded in the *bai* operon [41]. The products of this 7α -dehydroxylation pathway, 308 deoxycholate and lithocholate, are known hepatotoxins associated with liver cancer [42] and 309 hypothesized to also promote CRC [43]. Although intensely studied at a mechanistic level, these 310 factors are not (well) represented in general databases that can be used for metagenome annotation 311 (Supplementary Fig. 5). Thus, we built a targeted metagenome annotation workflow based on

312 Hidden Markov Models to identify and quantify virulence factors and toxicity pathways of interest in 313 CRC. Additionally, we used co-abundance clustering to infer operon completeness for factors encoded 314 by multiple genes (Methods, Extended Data 9, Supplementary Fig. 5). While fadA, bft, the pks island 315 and the bai operon were clearly detectable in deeply sequenced fecal metagenomes, they varied 316 broadly with respect to abundance, significance and cross-study consistency of enrichment (Fig. 4c): 317 fadA and pks were significantly enriched in CRC metagenomes (P = 5.3E-10 and 4.1E-4 respectively), 318 whereas no significant abundance difference could be detected for bft in fecal metagenomes, despite 319 reports on its enrichment in the mucosa of CRC patients [44], its carcinogenic effect in mouse models 320 [45], and synergistic action with pks [46]. Our quantification of the bai operon showed a highly 321 significant enrichment in CRC metagenomes (P = 1.6E-9) observed across all five studies (Fig. 4d) at 322 an average abundance that exceeded fadA and pks copy numbers (Fig. 4c). Metagenome analysis 323 indicated that at least four Clostridiales species (including the well characterized C. scindens and C. 324 hylemonae [47, 48]) have a (near) complete 7α -dehydroxylation pathway contributing to the observed 325 enrichment of bai operon copies (Extended Data 9). To validate this finding and further explore its 326 value towards diagnostic application, we developed a targeted quantification assay for the baiF gene 327 based on quantitative PCR (gPCR, see Methods). Quantification of baiF by gPCR using genomic DNA 328 from 47 fecal samples of the DE study population was found to be similar to, yet more sensitive than 329 by metagenomics (Fig. 4e). Gut microbial baiF copy numbers clearly distinguished CRC patients from 330 controls (P = 0.001) at an AUROC of 0.77, which in this subset of samples is surpassed by only a 331 single species marker for CRC (Extended Data 9). Although consistent with increased deoxycholate 332 metabolite levels reported for serum and stool samples of CRC patients [49], this finding does not 333 imply 7a-dehydroxylation pathway activity. We therefore quantified baiF expression using RNA 334 extracts from the same set of fecal samples, and found also transcript levels to be elevated in CRC 335 patients (Fig. 4f). The observed weak correlation of baiF expression with genomic abundance (Fig. 4f) 336 might be explained by dynamic transcriptional regulation [47] and bai expression in feces might not 337 accurately reflect the tumor microenvironment. Taken together, these data suggest gut microbial 338 metabolic markers to be meaningful and highly predictive of CRC status.

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340 Validation of CRC signatures in independent study populations

341 Even though CRC classification accuracy for both species and functions were evaluated on 342 independent data, we nonetheless sought to confirm it using two additional study populations from 343 Italy (IT1 and IT2, combined N = 61 CRC, N = 62 CTR, [27], see Methods, Table 1) and one from 344 Japan (JP, N = 40 CRC, N = 40 CTR, see Methods, Table 1). The overlap of single species 345 associations detected in the IT2 study and those from the meta-analysis was found to vary within the 346 range seen for the other studies, whereas for IT1 and JP the overlap was slightly lower (cf. study 347 precision in Supplementary Fig. 2, Extended Data 10). Nonetheless, the AUROC of LOSO 348 classification models based on species ranged between 0.79 and 0.81 and that for the classifiers 349 based on eggNOG from 0.71 to 0.92 (Fig. 5ab). We also validated CRC enrichment of fadA, pks and 350 bai genes in these three study populations (Fig. 5c). Altogether these results highlight consistent 351 alterations in the gut microbiome of CRC patients across eight study populations from seven countries 352 in three continents.

353

354 DISCUSSION

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356 Through extensive and statistically rigorous validation, in which data from studies used for training is 357 strictly separated from that for testing, our meta-analysis firmly establishes that gut microbial 358 signatures are highly predictive of CRC (see also [27]). In particular metagenomic classifiers trained 359 on species profiles from multiple studies maintained an AUROC of at least 0.8 in seven out of eight 360 data sets and achieved an accuracy similar to the fecal occult blood test, a standard non-invasive 361 clinical test for CRC (Supplementary Fig. 6, cf. [8]). These results thus suggest that polymicrobial 362 CRC classifiers are globally applicable and can overcome technical and geographical study 363 differences, which we found to generally impact observed microbiome composition more than the 364 disease itself (Fig. 1c, Extended Data 1, 2). The generalization accuracy of classifiers across studies 365 seen here is higher than that reported in 16S rRNA gene amplicon sequencing studies, which are 366 characterized by even larger heterogeneity across studies [16, 18] (Supplementary Fig. 7).

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Previous microbiome meta-analyses suggested that the majority of gut microbial taxa differing in any given case-control study reflect general dysbiosis rather than disease-specific alterations illustrating the difficulty of establishing disease-specific microbiome signatures [17, 19]. Here, by combining data across studies for training (LOSO), we were able to develop disease-specific signatures that maintained false positive control on diabetes and IBD metagenomes at a very similar level as for CRC (**Fig. 3c**) despite these diseases having shared effects on the gut microbiome [17, 50] and an increased comorbidity risk [51].

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376 Although for diagnostic purposes, unresolved causality between microbial and host processes during 377 CRC development are not a central concern, elucidating the underlying mechanisms would greatly 378 enhance our understanding of colorectal tumorigenesis. Towards this goal, we developed both broad 379 and targeted annotation workflows for functional metagenome analysis. First, we found functional 380 signatures based on the abundances of orthologous groups of microbial genes to yield accuracies as 381 high as taxonomic signatures (Fig. 3), which raises the hope for future improvements in metagenome 382 annotation to translate into microbiome signature refinements. Second, by investigating potentially 383 carcinogenic bacterial virulence and toxicity mechanisms taking a targeted metagenome annotation 384 approach, we confirmed highly significant enrichments of the colibactin-producing pks gene cluster 385 and the Fusobacterium nucleatum adhesin FadA in CRC metagenomes (Fig. 4c). Our results support 386 the clinical relevance of these factors adding to the experimental evidence for their carcinogenic 387 potential [46, 52-54]. We further examined the bai operon, encoding enzymes that produce secondary 388 bile acids via 7α-dehydroxylation, as an example of toxic host-microbial co-metabolism (see [27] for 389 another intriguing example). While α -dehydroxylated bile acids are established liver carcinogens [42]. 390 their contribution to CRC is less clear [43]. Here, we have, for the first time, shown bai to be highly 391 enriched in stool from CRC patients (Fig. 4cd) and confirmed this finding at both the genomic and the 392 transcriptomic level using qPCR (Fig. 4ef). As bai enrichment (and expression) is likely a 393 consequence of a diet rich in fat and meat [55], it is intriguing to explore whether bai could be used as

394 a surrogate microbiome marker for such difficult-to-measure dietary CRC risk factors. To further 395 unravel the molecular underpinning of these dietary CRC risk factors, molecular pathological 396 epidemiology studies that investigate the mucosal microbiome as part of the tumor microenvironment, 397 hold great potential [56, 57]. However, they will require more comprehensive diet questionnaires, 398 medical records, and molecular tumor characterizations than are available for the study populations 399 analyzed here. In this context, carcinogens possibly contained in the virome also warrant further 400 investigation [58, 59], but for this goal, metagenomic data needs to be generated with protocols 401 optimized for virus enrichment [60].

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403 Taken together, our results and those by Thomas, Manghi et al. [27], strongly support the promise of 404 microbiome-based CRC diagnostics. Both taxonomic and metabolic gut microbial marker genes 405 established in these meta-analyses could form the basis of future diagnostic assays that are 406 sufficiently robust, sensitive, and cost-effective for clinical application. The targeted gPCR-based 407 quantification of the baiF gene is a first step in this direction. Our metagenomic analysis of this and 408 other virulence and toxicity markers bridge to existing mechanistic work in preclinical models and 409 could enable future work aiming to precisely determine the contribution of gut microbiota to CRC 410 development.

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413 Data and Code Availability

The raw sequencing data for the samples in the DE study that had not been published before (see Methods), are made available in the European Nucleotide Archive (ENA) under the study identifier PRJEB27928. Metadata for these samples are available as **Supplementary Table S6**.

417 For the other studies included here, the raw sequencing data can be found under the following ENA

418 identifiers: PRJEB10878 for [11], PRJEB12449 for [10], ERP008729 for [9], and ERP005534 for [8].

The independent validation cohorts can be found in SRA under the identifier SRP136711 for [27] and in the DDBJ database under the ID DRA006684.

- 421 Filtered taxonomic and functional profiles used as input for the statistical modeling pipeline are
- 422 available in **Supplementary Data 1**.

423 The code and all analysis results can be found under https://github.com/zellerlab/crc_meta.

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450

451 **Competing Interest**

452 P. Bork, G. Zeller, A.Y. Voigt, and S. Sunagawa are named inventors on a patent (EP2955232A1:
453 Method for diagnosing colorectal cancer based on analyzing the gut microbiome).

454

455 Author Contributions

456 G.Z., M.A., P.B. conceived and supervised the study. P.S.-K., N.H., C.M.U., H.B., E.V., R.S. recruited 457 patients and collected samples. E.K., A.Y.V., S.Sunagawa, P.B. generated metagenomic data. A.M., 458 P.T.P., J.S.F., A.P., S.Sunagawa, L.P.C., G.Z., M.A. developed metagenomic profiling workflows 459 and/or performed taxonomic and functional profiling. J.W., G.Z., K.Z., P.T.P., A.K., M.A., N.S. 460 performed statistical analysis and/or developed statistical analysis workflows. E.K. and R.P.P. 461 designed and performed validation experiments, A.M.T., P.M., S.G., D.S., S.M., H.S., S.Shiba, T.S., 462 S.Y., T.Y., L.W., A.N., N.S. provided additional validation data, J.W., G.Z., M.A., P.T.P., P.B. designed figures. G.Z., J.W., M.A., P.B., wrote the manuscript with contributions from P.T.P., A.M., 463 464 S.Sunagawa, L.P.C., E.K., A.Y.V., E.V., R.S., P.S.K., H.B., E.N., N.S., L.W. All authors discussed and 465 approved the manuscript.

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467

468 Figure Captions

469

470 Figure 1. Despite study differences, meta-analysis identifies a core set of gut microbes471 strongly associated with CRC.

(a) Meta-analysis significance of gut microbial species derived from blocked Wilcoxon tests (n=574
independent observations) is given by bar height (false discovery rate, FDR, of 0.05). (b) Underneath,
species-level significance as computed by two-sided Wilcoxon test (FDR-corrected P-value) and
generalized fold change (Methods) within individual studies are displayed as heatmaps in gray and

476 color, respectively (see color bars and Table 1 for details on studies included). Species are ordered by 477 meta-analysis significance and direction of change. (c) For a core of highly significant species (meta-478 analysis FDR 1E-5), association strength is quantified by the area under the Receiver Operating 479 Characteristics curve (AUROC) across individual studies (color coded diamonds) and 95% confidence 480 intervals are indicated by gray lines. Family-level taxonomic information is color-coded above species 481 names (numbers in brackets are mOTU species identifiers, see Methods). (d) Variance explained by 482 disease status (CRC vs controls) is plotted against variance explained by study effects for individual 483 microbial species with dot size proportional to abundance (Methods); core microbial markers are 484 highlighted in red. F. nucleatum – Fusobacterium nucleatum.

485

Figure 2. Co-occurrence analysis of CRC-associated gut microbial species reveals four clusters preferentially linked to specific patient subgroups.

488 (a) The heatmap shows for all CRC patients (n=285 independent samples) if the respective sample is 489 positive for each of the core set of microbial marker species (see Methods for adjustment of positivity 490 threshold). Samples are ordered according to the sum of positive markers and marker species are 491 clustered based on Jaccard similarity of positive samples, resulting in four clusters (Methods). Barplots 492 in (b), (c), and (d) show the fraction of CRC samples that are positive for marker species clusters 493 (defined as the union of positive marker species) broken down by patient subgroups based on 494 differences in tumor location, sex, or CRC stage, respectively. Statistically significant associations 495 between CRC subgroups and marker species clusters were identified using the Cochran-Mantel-496 Haenszel test blocked for study effects and are indicated above bars (P < 0.1).

497

Figure 3. Both taxonomic and functional metagenomic classification models generalize across studies in particular when trained on data from multiple studies.

500 CRC classification accuracy resulting from cross validation within each study (gray boxes along 501 diagonal) and study-to-study model transfer (external validations off diagonal) as measured by 502 AUROC for classifiers trained on (a) species and (d) eggNOG gene family abundance profiles. The 503 last column depicts the average AUROC across external validations. Classification accuracy, as 504 evaluated by AUROC on a held-out study, improves if taxonomic (b) or functional (e) data from all 505 other studies are combined for training (leave-one-study-out, LOSO validation) relative to models 506 trained on data from a single study (study-to-study transfer, average and standard deviation shown). 507 Bar height for study-to-study transfer corresponds to the average of four classifiers (error bars indicate 508 standard deviation, n=4). (c) Combining training data across studies substantially improves CRC 509 specificity of the (LOSO) classification models relative to models trained on data from a single study 510 (depicted by bar color, as in (c) and (d)) as assessed by the false positive rate (FPR) on fecal samples 511 from patients with other conditions (see legend). Bar height for study-to-study transfer corresponds to 512 the average FPR across classifiers (n=5) with error bars indicating the standard deviation of FPR 513 values observed.

514

515 Figure 4. Meta-analysis identifies consistent functional changes in CRC metagenomes.

516 (a) Meta-analysis significance of gut metabolic modules derived from blocked Wilcoxon tests (n=574 517 independent samples) is indicated by bar height (top panel, FDR of 0.01). Underneath, the 518 generalized fold change (Methods) for gut metabolic modules [31] within individual studies is displayed 519 as heatmap (see color key below (b)). Metabolic modules are ordered by significance and direction of 520 change. A higher-level classification of the modules is color-coded below the heatmap for the four 521 most common categories (colors as in (b), white indicating other classes). (b) Normalized log 522 abundances for these selected functional categories is compared between controls (CTR) and 523 colorectal cancer cases (CRC). Abundances are summarized as geometric mean of all modules in the 524 respective category and statistical significance determined using blocked Wilcoxon tests (n=574 525 independent samples, see Methods). (c) Normalized log abundances for virulence factors and toxins 526 compared between metagenomes of controls (CTR) and colorectal cancer cases (CRC) (significant 527 differences P < 0.05 were determined by blocked Wilcoxon test, n=574 independent samples, see 528 Methods for gene identification and quantification in metagenomes; fadA: gene encoding 529 Fusobacterium nucleatum adhesion protein A, bft: gene encoding Bacteroides fragilis enterotoxin, pks: 530 genomic island in *Escherichia coli* encoding enzymes for the production of genotoxic colibactin, and 531 bai: bile acid inducible operon present in some Clostridiales species encoding bile acid converting 532 enzymes). (d) Meta-analysis significance (uncorrected P-value) as determined by blocked Wilcoxon 533 tests (n=574 independent samples) and generalized fold change within individual studies are 534 displayed as bars and heatmap, respectively, for the genes contained in the bai operon. Due to high 535 sequence similarity to baiF, baiK was not independently detectable with our approach. (e) 536 Metagenomic quantification of baiF (metag. ab. - normalized relative abundance) is plotted against 537 qPCR quantification in genomic DNA (gDNA) extracted from a subset of DE samples (n=47), with 538 Pearson correlation (r) indicated (see Methods). (f) Expression of baiF determined via qPCR on 539 reverse-transcribed RNA from the same samples in contrast to genomic DNA (as in e). The boxplots 540 on the side of (e), (f) show the difference between cancer (CRC) and control (CTR) samples in the 541 respective qPCR quantification (P-values on top were computed using a one-sided Wilcoxon test). All 542 boxplots show interguartile ranges (IQR) as boxes with the median as a black horizontal line and 543 whiskers extending up to the most extreme points within 1.5-fold IQR.

544

545 Figure 5. Meta-analysis results are validated in three independent study populations

546 CRC classification accuracy for independent datasets, two from Italy and one from Japan (see 547 Supplementary Table S2), is indicated by bar height for single study (white) and leave-one-study-out 548 (grey) models using either (a) species or (b) eggNOG gene family abundance profiles (cf. Fig. 3). Bar 549 height for single study models corresponds to the average of five classifiers (error bars indicate 550 standard deviation, n=5). (c) Normalized log abundances for virulence factors and toxins (cf. Figure 551 4c) compared between controls (CTR) and colorectal cancer cases (CRC). P-values were determined 552 by blocked, one-sided Wilcoxon tests (n=193 independent samples). Boxes represent interguartile 553 ranges (IQR) with the median as a black horizontal line and whiskers extending up to the most 554 extreme points within 1.5-fold IQR.

555

556

557 Table 1. Fecal metagenomic studies of colorectal cancer included in this meta-analysis.

558 See Methods for inclusion criteria and **Supplementary Table S2** for extended meta-data. For a 559 detailed description of patient recruitment and data generation for the DE study, see Methods. The 560 data for 38 samples from the DE study had been published previously as part of an independent 561 validation cohort in [8].

Country Code	Reference	No. of cases	No. of controls		
FR	Zeller et al., 2014 [8]	53	61		
AT	Feng et al., 2015 [9]	46	63		
CN	Yu et al., 2017 [11]	74	54		
US	Vogtmann et al., 2016 [10]	52	52		
DE	this study	60	60		
External validation cohorts					
IT1	[27]	29	24		
IT2	[27]	32	28		
JP	Courtesy of T. Yamada et al.	40	40		

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565 References

- 566 1. Tringe, S.G. and E.M. Rubin, *Metagenomics: DNA sequencing of environmental samples.* 567 Nat. Rev. Genet., 2005. **6**(11): p. 805-814.
- 568 2. Tremaroli, V. and F. Bäckhed, *Functional interactions between the gut microbiota and host metabolism.* Nature, 2012. **489**(7415): p. 242-249.
- 570 3. Lynch, S.V. and O. Pedersen, *The Human Intestinal Microbiome in Health and Disease*. N. Engl. J. Med., 2016. **375**(24): p. 2369-2379.
- 4. Qin, J., et al., A metagenome-wide association study of gut microbiota in type 2 diabetes.
 Nature, 2012. 490(7418): p. 55-60.
- 574 5. Karlsson, F.H., et al., *Gut metagenome in European women with normal, impaired and diabetic glucose control.* Nature, 2013. **498**(7452): p. 99-103.
- 576 6. Qin, J., et al., *A human gut microbial gene catalogue established by metagenomic sequencing.* Nature, 2010. **464**(7285): p. 59-65.
- 578 7. Schirmer, M., et al., *Dynamics of metatranscription in the inflammatory bowel disease gut microbiome*. Nat Microbiol, 2018. 3(3): p. 337-346.
- 580 8. Zeller, G., et al., *Potential of fecal microbiota for early-stage detection of colorectal cancer.*581 Mol. Syst. Biol., 2014. **10**(11): p. 766.
- 582 9. Feng, Q., et al., *Gut microbiome development along the colorectal adenoma-carcinoma sequence.* Nat. Commun., 2015. 6: p. 6528.
- 58410.Vogtmann, E., et al., Colorectal Cancer and the Human Gut Microbiome: Reproducibility with585Whole-Genome Shotgun Sequencing. PLoS One, 2016. 11(5): p. e0155362.

586	11.	Yu, J., et al., Metagenomic analysis of faecal microbiome as a tool towards targeted non-
587		invasive biomarkers for colorectal cancer. Gut, 2017. 66 (1): p. 70-78.
588	12.	Bedarf, J.R., et al., Functional implications of microbial and viral gut metagenome changes in
589		early stage L-DOPA-naïve Parkinson's disease patients. Genome Med., 2017. 9(1): p. 39.
590	13.	Schmidt, T.S.B., J. Raes, and P. Bork, The Human Gut Microbiome: From Association to
591		Modulation. Cell. 2018. 172(6); p. 1198-1215.
592	14.	Forslund, K., et al., Disentangling type 2 diabetes and metformin treatment signatures in the
593		human out microbiota Nature 2015 528 (7581) p. 262-266
594	15	Costea PL et al. Towards standards for burnan fecal sample processing in metagenomic
595	10.	studies Nat Biotechnol 2017 35(11): n 1069-1076
596	16	Lozunne CA et al. Meta-analysis of studies of the human microbiota. Genome Res
597	10.	2013 23 (10): n 1704-1714
508	17	Duvallet C at al Meta Analysis Of Microbiome Studies Identifies Shared And Disease-
500	17.	Sherific Patterns 2017
600	18	Shah $M \leq a_1 \leq 1 \leq r_2$
601	10.	Sitial, M.S., et al., Levelaging sequence-based nated infinitional community survey data to
602	10	Baselli E, et al. Mashina Learning Mata analysis of Large Matagenetic Datasets: Tools and
602	19.	Pasoni, L., et al., Machine Learning Meta-analysis of Large Metageronic Datasets. Tools and Biological Insisters Dios Computer Biol. 2016, 12 (7): p. a1004077
604	20	Diological integrities, FLOS Complete Bioli, 2010, 12(1), P. 61004911.
60F	20.	Dat, Z., et al., Multi-conort analysis of conortectal cancer interagenome interaction and entered
605	01	Dacteria across populations and universal bacteria markers. Microbione, 2016. 0(1), p. 70.
600	21.	
607	00	333(7098), p. 023-028.
600	ZZ.	manese, A., et al., <i>Microbial abundance, activity, and population genomic proming with</i>
609	00	<i>MOTOS</i> . Nature communications, 2019. formany accepted for publication.
010	23.	Ruitina, J.R., et al., MOCAT2. a metagenomic assembly, annotation and proming framework.
611	04	Bioinformatics, 2016. 32 (16): p. 2520-2523.
612	24.	Hotnorn, T., et al., A Lego System for Conditional Inference. Am. Stat., 2006. 60(3): p. 257-
613	05	
614	25.	Mandal, S., et al., Analysis of composition of microbiomes: a novel method for studying
615	~~	microbial composition. Microb Ecol Health Dis, 2015. 26: p. 27663.
616	26.	I Jalsma, H., et al., A bacterial driver-passenger model for colorectal cancer: beyond the usual
617		suspects. Nat Rev Microbiol, 2012. 10 (8): p. 575-82.
618	27.	I homas, A.M., et al., Metagenomic analysis of colorectal cancer datasets identifies cross-
619		cohort microbial diagnostic signatures and a link with choline degradation. co-submitted to
620		Nature Medicine, 2018.
621	28.	Huerta-Cepas, J., et al., eggNOG 4.5. a hierarchical orthology framework with improved
622		functional annotations for eukaryotic, prokaryotic and viral sequences. Nucleic Acids Res.,
623		2016. 44 (D1): p. D286-93.
624	29.	Kanehisa, M., et al., Data, information, knowledge and principle: back to metabolism in KEGG.
625		Nucleic Acids Res., 2014. 42(Database issue): p. D199-205.
626	30.	Li, J., et al., An integrated catalog of reference genes in the human gut microbiome. Nat.
627		Biotechnol., 2014. 32 (8): p. 834-841.
628	31.	Vieira-Silva, S., et al., Species-function relationships shape ecological properties of the human
629		<i>gut microbiome</i> . Nat Microbiol, 2016. 1 (8): p. 16088.
630	32.	Hirayama, A., et al., Quantitative metabolome profiling of colon and stomach cancer
631		microenvironment by capillary electrophoresis time-of-flight mass spectrometry. Cancer Res,
632		2009. 69 (11): p. 4918-25.
633	33.	Denkert, C., et al., Metabolite profiling of human colon carcinomaderegulation of TCA cycle
634		and amino acid turnover. Mol Cancer, 2008. 7: p. 72.
635	34.	Mal, M., et al., Metabotyping of human colorectal cancer using two-dimensional gas
636		chromatography mass spectrometry. Anal Bioanal Chem, 2012. 403(2): p. 483-93.
637	35.	Weir, T.L., et al., Stool microbiome and metabolome differences between colorectal cancer
638		patients and healthy adults. PLoS One, 2013. 8(8): p. e70803.
639	36.	Goedert, J.J., et al., Fecal metabolomics: assay performance and association with colorectal
640		<i>cancer.</i> Carcinogenesis, 2014. 35 (9): p. 2089-2096.
641	37.	Aykan, N.F., <i>Red meat and colorectal cancer.</i> Oncology Reviews, 2015. 9(1).
642	38.	World Cancer Research Fund / American Institute for Cancer Research, Diet, Nutrition,
643		Physical Activity and Cancer: a Global Perspective, in Continuous Update Project Expert
644		Report. 2018.
645	39.	Dutilh, B.E., et al., Screening metatranscriptomes for toxin genes as functional drivers of
646		human colorectal cancer. Best Pract Res Clin Gastroenterol, 2013. 27(1): p. 85-99.

647 648	40.	Sears, C.L. and W.S. Garrett, <i>Microbes, Microbiota, and Colon Cancer.</i> Cell Host Microbe, 2014. 15 (3): p. 317-328.
649	41.	Ridlon, J.M., et al., Consequences of bile salt biotransformations by intestinal bacteria. Gut
650		Microbes, 2016. 7 (1): p. 22-39.
651 652	42.	Yoshimoto, S., et al., Obesity-induced gut microbial metabolite promotes liver cancer through senescence secretome. Nature, 2013. 499 (7456): p. 97-101.
653	43.	Ajouz, H., D. Mukherji, and A. Shamseddine, Secondary bile acids: an underrecognized cause
654		of colon cancer. World Journal of Surgical Oncology, 2014. 12(1): p. 164.
655	44.	Boleij, A., et al., The Bacteroides fragilis toxin gene is prevalent in the colon mucosa of
656		colorectal cancer patients. Clin. Infect. Dis., 2015. 60(2): p. 208-215.
657	45.	Wu, S., et al., A human colonic commensal promotes colon tumorigenesis via activation of T
658		helper type 17 T cell responses. Nat. Med., 2009. 15(9): p. 1016-1022.
659	46.	Dejea, C.M., et al., Patients with familial adenomatous polyposis harbor colonic biofilms
660		containing tumorigenic bacteria. Science, 2018. 359 (6375): p. 592-597.
661	47.	Ridlon, J.M., D.J. Kang, and P.B. Hylemon, Isolation and characterization of a bile acid
662		inducible 7alpha-dehydroxylating operon in Clostridium hylemonae TN271. Anaerobe, 2010.
663		16 (2): p. 137-46.
664	48.	Mallonee, D.H., W.B. White, and P.B. Hylemon, Cloning and sequencing of a bile acid-
665		inducible operon from Eubacterium sp. strain VPI 12708. Journal of Bacteriology, 1990.
666		172 (12): p. 7011-7019.
667	49.	Ocvirk, S. and S.J.D. O'Keefe, Influence of Bile Acids on Colorectal Cancer Risk: Potential
668		Mechanisms Mediated by Diet-Gut Microbiota Interactions. Curr. Nutr. Rep., 2017. 6(4): p.
669		315-322.
670	50.	Gevers, D., et al., The treatment-naive microbiome in new-onset Crohn's disease. Cell Host
6/1	- 4	Microbe, 2014. 15 (3): p. 382-392.
672	51.	Viennot, S., et al., Colon cancer in inflammatory bowel disease: recent trends, questions and
6/3	50	answers. Gastroenterol. Clin. Biol., 2009. 33 Suppl 3: p. S190-201.
674	52.	Rubinstein, M.R., et al., Fusobacterium nucleatum Promotes Colorectal Carcinogenesis by
675		Modulating E-Cadherin/B-Catenin Signaling via its FadA Adhesin. Cell Host Microbe, 2013.
670	52	14(2). β. 195-200. Kestia Δ.D. et al. Eucohostarium nucleatum natantiatas intestinal tumorizanasis and
679	53.	Kostic, A.D., et al., Fusopacterium nucleatum potentiales intestinal tumongenesis and medulates the tumor immune microenvironment. Cell Hest Microbe, 2012, 14(2): p. 207, 215
670	54	Arthur I.C. at al. Intestinal inflammation targets cancer inducing activity of the microhiota
680	54.	Arthur, J.C., et al., intestinal initiation largets cancer-inducing activity of the microbiola.
681	55	Buddy B.S. Diet and excretion of hile acids. Cancer Res. 1081. /1 /0 Dt 2): p. 3766.8
682	56	Ogino S, et al. Integrative analysis of exogenous endogenous tumour and immune factors
683	50.	for precision medicine Gut 2018 67(6): p. 1168-1180
684	57	Onino S et al. Molecular pathological enidemiology of colorectal neoplasia: an emerging
685	57.	transdisciplinary and interdisciplinary field Gut 2011 60(3): p. 307-411
686	58	Hannigan G.D. et al. Diagnostic Potential and Interactive Dynamics of the Colorectal Cancer
687	00.	Virome MBio 2018 9(6)
688	59	zur Hausen H. Red meat consumption and cancer: reasons to suspect involvement of bovine
689	00.	infectious factors in colorectal cancer. Int J Cancer. 2012 130 (11): p. 2475-83
690	60	Shkoporov A N et al. Reproducible protocols for metagenomic analysis of human faecal
691	00.	phageomes. Microbiome. 2018. 6(1): p. 68.
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- 695 Methods
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697 Study inclusion and data acquisition

We used PubMed to search for studies that published fecal shotgun metagenomic data of human colorectal cancer patients and healthy controls. The search term, all hits, and the justification for exclusion or inclusion are available in **Supplementary Table S1**. Raw fastq files were downloaded for the four included studies from the European Nucleotide Archive, using the following ENA identifiers: PRJEB10878 for [11], PRJEB12449 for [10], ERP008729 for [9], and ERP005534 for [8].

703

704 DE study recruitment and sequencing

705 The German (DE) study population data consist of 60 fecal CRC metagenomes, 38 of which were 706 sequenced and published in [8] under ENA accession ERP005534. The fecal metagenomes from 707 additional 22 CRC patients recruited for the same ColoCare study (DKFZ, Heidelberg, [61, 62]) were 708 sequenced later as part of this work. All fecal samples were collected after colonoscopy. Sixty gender-709 and age-matched participants of the PRÄVENT study run by the same clinical investigators were 710 included as healthy controls; as these were not subjected to colonoscopy, the presence of 711 undiagnosed colorectal carcinomas cannot be completely ruled out but is expected to be unlikely due 712 to low prevalence of preclinical CRC in the general population [63].

713 Written informed consent was obtained from all additional 22 CRC patients and 60 controls. The study 714 protocol was approved by the institutional review board (EMBL Bioethics Internal Advisory Board) and 715 the ethics committee of the Medical Faculty at the University of Heidelberg. The study is in agreement 716 with the WMA Declaration of Helsinki and the Department of Health and Human Services Belmont 717 Report.

Genomic DNA was extracted from the fecal samples (preserved in RNALater) and libraries were
prepared as previously described [8]. Whole-genome shotgun sequencing was performed by using
Illumina HiSeq 2000 / 2500 / 4000 (Illumina, San Diego, USA) platforms at the Genomics Core
Facility, European Molecular Biology Laboratory, Heidelberg.

722

723 Independent validation cohorts

During the revision of this manuscript, we included three independent study populations for external validation. Two of them were recruited in Italy (IT1 and IT2) with informed consent from all participants and ethical approval by the Ethics committee of Azienda Ospedaliera of Alessandria and that of the European Institute of Oncology of Milan. Shotgun fecal metagenomic data was generated as described in [27].

The third study population was recruited in Japan (JP) with informed consent and ethical approval of the institutional review boards of the National Cancer Center Japan - Research Institute and the Tokyo Institute of Technology. DNA was extracted from frozen fecal samples using a GNOME DNA Isolation Kit (MP Biomedicals, Santa Ana, CA) with an additional bead-beating step as previously described [64]. DNA quality was assessed with an Agilent 4200 TapeStation (Agilent Technologies, Santa Clara CA). After final precipitation, the DNA samples were resuspended in TE buffer and stored at -80°C before further analysis. Sequencing libraries were generated with the Nextera XT DNA Sample Preparation Kit (Illumina, San Diego, CA). Library quality was confirmed with an Agilent 4200
TapeStation. Whole-genome shotgun sequencing was carried out on the HiSeq2500 platform
(Illumina). All samples were paired-end sequenced with a 150-bp read length to a targeted data set
size of 5.0 Gb.

740

741 Taxonomic profiling and data preprocessing

742 The metagenomic samples were quality controlled using MOCAT2's -rtf procedure, which is based on 743 the 'solexaqa' algorithm [23]. In particular, reads that map with at least 95% sequence identity and 744 alignment length of at least 45 bp to the human genome hg19 were removed. In a second step, 745 taxonomic profiles were generated with the mOTU profiler version 2.0.0 ([22, 65, 66] - see motu-746 tool.org and GitHub version tag 2.0.0) using the following parameters: -1 75, -g 2 and -c. Briefly, this 747 profiler is based on ten universal single-copy marker-gene families (COG0012, COG0016, COG0018, 748 COG0172, COG0215, COG0495, COG0525, COG0533, COG0541 and COG0552) [66]. These 749 marker-genes were extracted from >25,000 reference genomes and >3,000 metagenomic samples 750 allowing to profile prokaryotic species with a sequenced reference genome (ref-mOTUs) and ones 751 without (meta-mOTUs). The read count for a mOTU was calculated as median of the read count of the 752 genes that belonged to that mOTU.

mOTU profiles were first converted to relative abundances to account for library size. Then, profiles were filtered to focus on a set of species that are confidently detectable in multiple studies. Specifically, microbial species that did not exceed a maximum relative abundance of 1E-03 in at least of the studies were excluded from further analysis, together with the fraction of unmapped metagenomic reads.

758

759 Functional metagenome profiling and data preprocessing

760 High-guality reads (same guality filtering as for taxonomic profiling) were aligned against a combined 761 database (IGChg38 hereafter) consisting of the hg38 release of the human reference genome and the 762 integrated gene catalog (IGC) containing 9.9 million non-redundant microbial genes [30] using BWA 763 mem [67] (Version: 0.7.15-r1140) with default parameters. The purpose of adding the human genome 764 to the reference database was to filter out reads that mapped as well or better to some human 765 sequence than to any bacterial gene. Alignments were computed separately for paired-end and single 766 read libraries (single reads could result from read pairs where one read was filtered out in the quality 767 filtering procedure described above). Alignments were then filtered to only retain those longer than 768 50bp with >95% sequence identity. Then the highest scoring alignment(s) was/were kept for each 769 read. As IGChg38 is a database of predominantly genes and not genomes, there will be a substantial 770 proportion of read-pairs where one end maps within the gene while the other end does not - it either 771 maps to an adjacent gene or remains unmapped due to intergenic regions not contained in the 772 database. Therefore, we counted a whole read-pair aligning to a gene when (i) both ends from a read 773 pair map to the same gene, (ii) only one end from a read-pair maps to the gene, or (iii) a read from the 774 single read library maps to the gene. We then counted only the read-pairs that map uniquely to one 775 gene in the IGC, thus excluding ambiguous read pairs mapping with similarly high scores to multiple 776 genes in the database. For a given metagenomic sample, we further normalized the abundance of each IGC gene by the length of that gene. We then estimated relative abundance of IGC genes by
dividing gene abundances by the total abundance of all genes in IGC (excluding the human
chromosomes).

Because metagenomes from CRC patients were not included when the IGC was constructed, we analyzed how well CRC-associated species as identified in this meta-analysis were represented in the IGC. Using a phylogenetic marker gene (COG0533), which is also used by the species profiling workflow on which the meta-analysis is based, for 24 out of the 29 core CRC-associated species we found a match in the IGC with at least 90% nucleotide identity, indicating that a sequence from the same species (above 93.1% identity) or a slightly more distant relative is present in the IGC (**Supplementary Fig. 8**).

The relative abundance of eggNOG orthologous groups [28] was estimated by summing relative abundances of genes annotated to belong to the same eggNOG orthologous group as of the most recent annotations provided by MOCAT2 [23]. To obtain KEGG orthologous groups (KO) and pathway abundances, we applied the same procedure, but using KEGG annotations for IGC provided by MOCAT2 [29].

792

793 Overview over statistical analyses

For univariate association testing between the abundances of microbial taxa or gene functions we used nonparametric tests throughout; all of these were two-sided Wilcoxon tests except were otherwise noted. To account for potential confounding and heterogeneity between data sets we employed a stratified version of the Wilcoxon test [24] (see below for details). ANOVA was conducted on rank-transformed data. Significance of binary co-occurrence patterns was assessed using (stratified) Cochrane-Mantel-Haenszel tests.

800 Multivariable analysis was done with strict separation between training and test data. This importantly 801 also pertained to feature selection, which was either done via the LASSO [68] or by nested cross-802 validation procedures to avoid overoptimistic performance assessment [69] (see below for details). All 803 samples included in this meta-analysis came from distinct individuals to ensure that generalization 804 across subjects – rather than across timepoints within a given subject – is assessed.

805

806 Confounder analysis

807 To quantify the effect of potential confounding factors relative to that of CRC on single microbial 808 species, we used an ANOVA-type analysis. The total variance within the abundance of a given 809 microbial species was compared to the variance explained by disease status and the variance 810 explained by the confounding factor akin to a linear model including both CRC status and confounding 811 factor as explanatory variables for species abundance. Variance calculations were performed on ranks 812 in order to account for non-Gaussian distribution of microbiome abundance data. Potential 813 confounders with continuous values were transformed into categorical data either as guartiles or for 814 the case of body mass index (BMI) into lean/obese/overweight according to conventional cutoffs (lean: 815 < 25, obese: 25 - 30, overweight: > 30).

816

817 Univariate meta-analysis for the identification of CRC-associated gut microbial species

Significance of differential abundance was tested on a per-species basis using a blocked Wilcoxon test implemented in the R coin package [24]. Informed by the results of the preceding confounder analysis, we blocked for `study` and additionally `colonoscopy` in the CN study. Within this framework, significance is tested against a conditional null distribution derived from permutations of the observed data. Notably, permutations are performed within each block in order to control for variations in block size and composition. To adjust for multiple hypothesis testing, P-values were adjusted using the false-discovery rate (FDR) method [70].

825 As nonparametric effect size measures we used the area under the ROC curve (AUROC) with 826 permutation-based confidence intervals computed using the pROC package in R [71]. We further 827 developed a generalization of the (logarithmic) fold change that is widely used for other types of read 828 abundance data. This generalization is designed to have better resolution for sparse microbiome 829 profiles (where 0 entries can render median-based fold change estimates uninformative for the large 830 portion of species with a prevalence below 0.5). The generalized fold change (gFC) is computed as 831 mean difference in a set of pre-defined quantiles of the logarithmic CTR and CRC distributions (see 832 Extended Data 3 for further details; we used quantiles ranging from 0.1 to 0.9 in increments of 0.1).

For the retrospective analysis of study precision and recall for detecting microbial species associations from the meta-analysis, the true set was defined as the species which were associated at a given FDR in the meta-analysis. Then, we checked how well this set of species would be recovered using the single-study significance as determined by the Wilcoxon test. Study precision corresponds to the proportion of meta-analysis significant species among those detected as significant in a single study. Similarly, recall (or sensitivity) corresponds to the proportion of species out of the true set of metaanalysis significant species that were recovered in a given study.

840

841 Species co-occurrence and cluster analysis in CRC metagenomes

842 For the analysis of gut bacterial species co-occurring in CRC microbiomes, relative abundances of the 843 core set of associated species (excluding the CRC-depleted Clostridiales meta-mOTU [1296]) were 844 discretized into binary values to determine whether a CRC (metagenomic) sample is "positive" or 845 "negative" for a given microbial marker. To normalize for differences in prevalence (and therefore 846 specificity) of these markers we adjusted the threshold value, above which a sample is labeled 847 "positive" based on the abundance in healthy controls. For each microbial species, the 95th percentile 848 in healthy controls was used as threshold, which effectively results in adjusting the per-marker false 849 positive rate to 0.05. Based on the binarized species-by-sample matrix, species were then clustered 850 using the Jaccard dissimilarity as implemented in the vegan package in R [72]. Associations between 851 species clusters and meta-variables were tested as 2-by-n (where n is the number of categories in the 852 meta-variable tested) contingency tables using a Cochrane-Mantel-Haenszel test with study as 853 blocking factor as implemented in the coin package [24].

854

855 Multivariable statistical modeling workflow and model evaluation

As a main goal of our work is to assess the generalization accuracy of microbiome-based CRC classifiers across technical and geographic differences in patient populations, we extensively validated classification models across studies taking the following two approaches.

- In *study-to-study transfer* validation, metagenomic classifiers were trained on a single study and their performance externally assessed on all other studies (off-diagonal cells in **Fig. 3ac**). Effectively we implemented a nested cross validation procedure on the training study to compute within-study accuracy (cells on the diagonal in **Fig. 3ac**) and tune the model hyperparameters.
- In *leave-one-study-out* (LOSO) validation, data from one study was set aside as an external validation set, while the data from the remaining 4 studies was pooled as a training set on which we implemented the same nested cross validation procedure as for study-to-study transfer (see [19] for a more detailed description of LOSO).
- Bata preprocessing, model building, and model evaluation was performed using the SIAMCAT R
 package (<u>https://bioconductor.org/packages/SIAMCAT</u>, version <u>1.1.0</u>).
- 869

870 Preprocessing of taxonomic abundance profiles for statistical modeling

871 Relative abundances were first filtered to remove markers with low overall abundance and no variance 872 (an artifact for single-study data arising from the joint data filtering described above), log-transformed 873 (after adding a pseudo-count of 1E-05 to avoid non-finite values resulting from log(0), [73]) and finally 874 standardized as z-scores. Data were split into training and test set for 10 times repeated 10-fold 875 stratified cross validation (balancing class proportions across folds). For each split, a L1-regularized 876 (LASSO) logistic regression model [68] was trained on the training set, which was then used to predict 877 the test set. The lambda parameter, i.e. regularization strength was selected for each model to 878 maximize the area under the precision recall curve under the constraint that the model contained at 879 least 5 non-zero coefficients. Models were then evaluated by calculating the area under the Receiver 880 Operating Characteristics curve (AUROC) based on the posterior probability for the CRC class.

- In model transfer to a hold-out study, the holdout data were normalized for comparability in the same way as the training dataset by using the frozen normalization function in SIAMCAT, which retains the same features and re-uses the same normalization parameters (e.g. the mean of a feature for z-score standardization). Then, all 100 models derived from the cross validation on the training dataset (10 times repeated 10-fold CV) were applied to the holdout dataset and predictions were averaged across all models.
- In the LOSO setting, data from the four training studies were jointly processed as a single dataset in
 the same way as described above using 10 times repeated 10-fold stratified cross validation.
- 889

890 Preprocessing of functional abundance profiles

891 Functional profiles, such as eggNOG gene family or KEGG module abundance profiles were 892 preprocessed as described above for species profiles, but using 1E-06 as maximum abundance cutoff 893 and 1E-09 as a pseudo-count during log transformation. Since these abundance tables contained 894 several thousand input features we implemented an additional feature selection step, which was 895 nested properly into the cross-validation procedures as described above. This nested approach is 896 crucial to avoid over-optimistically biased performance estimates ([74], Chapter 7.10). Specifically, 897 features were filtered inside each training fold (without using any information from the test fold) by 898 selecting the 1600 features with highest single-feature AUROC values (for features depleted in CRC, 899 1 - AUROC was used for feature selection).

900

901 **Preprocessing of gene abundance profiles**

To ascertain the predictive power of a classifiers based on IGC gene abundances [30] we applied a series of filters to the abundance tables to reduce the number of genes that would be the input of the LASSO modelling. These filters where applied once on a per-study level and once in a leave-onestudy-out (LOSO) mode, where they were applied jointly to all studies in the training set, with the remaining one being held out for external validation.

- 907 The following filters were applied in this order:
- 908

All genes with 0 abundance in ≥15% of samples (regardless of CRC status) were discarded.

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 2. The remaining data was discretized using the equal frequencies method implemented in the
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- 912 3. As a feature selection procedure, mRMR (code version from 20 April 2009 downloaded from
 913 http://home.penglab.com/proj/mRMR/ on 3 Dec 2016) was run on the gene abundance table
 914 to retain the 100 top genes as output.

915 LASSO models were then built on log10-transformed abundances (pseudo-count of 10E-09, centered 916 and scaled) of the sets of 100 top genes returned by mRMR. The whole process was repeated 10 917 times in a 5-fold stratified cross-validation scheme to allow for an estimation of the confidence of the 918 AUROCs of the resulting models. We used the LiblineaR package (version 2.10-8) to build the LASSO 919 models in R and tested a sequence of 20 cost parameters (equivalent or the lambda parameter 920 controlling regularization strength) evenly spaced from 0.001² to 0.2². The cost parameter was 921 selected to maximize the AUROC within the training set.

922

923 External evaluation of disease-specificity of the metagenomic classifiers

To assess how disease-specific the predictions of the CRC models are, we applied these to data from case-control studies investigating other human diseases. Fecal metagenomic data of patients with Parkinson's disease [12], type 2 diabetes [4, 5], and inflammatory bowel disease [6, 7] were taxonomically profiled as described above. The parameters for quality control with MOCAT2 and for the mOTU profiler were the same as described above, except for the data from [6], where we used -I 50 (to set the threshold for minimum alignment length to 50) as the read length is shorter (average read length 71) compared to the other more recently generated Illumina shotgun metagenomic data.

931 Relative abundance data were treated exactly as another holdout dataset for each model, i.e. applying 932 the frozen normalization prediction routines as described above. For each CRC model applied to the 933 external datasets, a cutoff on its prediction output was adjusted to yield a false positive rate (FPR) of 934 0.1 on the controls of its respective (CRC) training set. Subsequently its FPR on metagenomes from 935 patients suffering from the above-mentioned (non-CRC) conditions was assessed to evaluate its 936 disease specificity. The rationale behind this is that a metagenomic classifier recognizing general 937 features of dysbiosis would be expected to predict CRC patients and those suffering from other 938 conditions at a similar rate; such a classifier would thus in the above-described evaluation display a 939 much higher FPR than on the controls of its training set. In contrast maintaining a low FPR in this

940 evaluation indicates that the classification model is based on CRC-specific features rather than941 hallmarks of general dysbiosis or nonspecific inflammation.

942

943 Functional profiling of gut metabolic modules (GMMs)

Gut metabolic modules were computed as originally proposed [31], using the KEGG KO profiles based on the IGC (see **Functional metagenome profiling** above) as input. Statistical analysis and generalized fold change calculations were performed analogously to species profiles (see above). Gut metabolic modules were summarized across functional groups (e.g. amino acid degradation) as geometric mean of all modules within the respective group.

949

950 Targeted functional analysis of virulence and toxicity pathways of potential relevance in CRC

951 To investigate toxins and virulence mechanisms that have previously been implicated with CRC [40]. 952 we constructed for each gene belonging to the respective virulence or toxicity pathway a hidden 953 Markov model (HMM). Each HMM was built from a multiple sequence alignment generated by 954 MUSCLE [76], containing the respective reference sequences and close homologs identified using 955 PSI-Blast [77]. Multiple sequence alignments are available together with the code for this paper 956 (https://github.com/zellerlab/crc_meta). Then, we screened the IGC metagenomic gene catalogue [30] 957 with each HMM using the HMMER software (version 3.1b2) [78]. Genes with an E-value below 1E-10 958 were filtered for uniqueness, since in some cases the HMMs would call different regions in the same 959 gene. For single gene virulence factors (i.e. fadA and bft), potential IGC hits were aligned against the 960 reference sequence using the Needleman-Wunsch algorithm in the EMBOSS package [79]. Hits were 961 then filtered based on percentage of sequence identity (cutoff: 40%) and sequence similarly to the 962 species relative abundance profiles based on maximum relative abundance (cutoff: 1E-07) in order to 963 exclude genes with limited relevance. Statistical analysis was performed on the sum of all genes.

964 For virulence pathways containing more than one gene, the IGC hits of each functional group within 965 the pathway were aligned against the respective reference sequence and filtered for percentage of 966 sequence identity and maximum abundance. Then, all hits were clustered based on the Pearson 967 correlation of the log-abundances across all samples using the Ward algorithm as implemented in the 968 hclust function in R. The gene clusters were filtered based on operon completeness (how many genes 969 of the operon were present in the cluster) and average correlation within the cluster (Extended Data 970 9). For statistical analysis, the genes in the selected gene clusters were summed up within each group 971 or all together for the overall analysis.

972

973 Quantitative PCR for baiF

974 Real-time quantitative PCR to quantify the abundance and expression of *baiF* was performed on a 975 subset of samples in the DE cohort (20 control and 24 colorectal cancer samples, see 976 **Supplementary Table S6**). For these samples, DNA and RNA extraction was done with the Allprep 977 PowerFecal DNA/RNA kit (Qiagen, Cat No: 80244) with additional RNAse and DNAse digestion steps, 978 respectively, as described by the manufacturer. DNA and RNA concentrations were determined by 979 Qubit Fluorometer (Invitrogen) and quality control of all RNA samples was done using an Agilent 2100 980 Bioanalyzer in combination with RNA 6000 Nano and Pico LabChip kits. 981 First-strand cDNA was synthesized by SuperScript IV VILO Master Mix with ezDNAse enzyme and 982 random hexamer primers (Invitrogen, catalogue number 11766500) as recommended by the 983 manufacturer. Reaction were performed as described in the protocol with one minor change of 984 temperature (incubation for the reverse transcription step at 55°C).

985 To quantify baiF relative to the total bacterial RNA/DNA in a sample, gPCR was performed in 986 triplicates for 16S rRNA and the baiF genes, using both cDNA and genomic DNA (gDNA) as template. 987 We used the following primers for baiF: TTCAGYTTCTACACCTG (forward), 988 GGTTRTCCATRCCGAACAGCG (reverse), and standard primers F515 and R806 for 16S [80]. RT-989 PCR reactions were prepared with a final primer concentration of 0.5 μ M, including 5 ng of genomic 990 DNA or 10 ng of cDNA in 20 μ l final reaction volume, and reactions were performed with SYBR Green gPCR mix on StepOne Real-Time PCR system (Thermo Fisler Scientific). Cycling conditions were as 991 follows; initial denaturation of 95°C for 10 min, then 40 cycles of denaturing at 95°C for 15 s, annealing 992 993 at 60°C for 60 s followed by melt curve analysis.

994 Delta-Ct values were calculated as difference between *baiF* and 16S Ct values. Significance of the 995 comparison between control and colorectal cancer samples was tested on the delta-Ct values using a 996 one-sided Wilcoxon test as a confirmation of metagenomic enrichment.

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998 Additional References

- 99961.Bohm, J., et al., Discovery of novel plasma proteins as biomarkers for the development of1000incisional hernias after midline incision in patients with colorectal cancer: The ColoCare study.1001Surgery, 2017. 161(3): p. 808-817.
- Liesenfeld, D.B., et al., Metabolomics and transcriptomics identify pathway differences
 between visceral and subcutaneous adipose tissue in colorectal cancer patients: the ColoCare
 study. Am J Clin Nutr, 2015. 102(2): p. 433-43.
- 100563.Pox, C.P., et al., Efficacy of a nationwide screening colonoscopy program for colorectal1006cancer. Gastroenterology, 2012. 142(7): p. 1460-7 e2.
- 100764.Furet, J.P., et al., Comparative assessment of human and farm animal faecal microbiota using1008real-time quantitative PCR. FEMS Microbiol Ecol, 2009. 68(3): p. 351-62.
- 100965.Mende, D.R., et al., Accurate and universal delineation of prokaryotic species. Nat. Methods,10102013. 10(9): p. 881-884.
- 1011 66. Sunagawa, S., et al., *Metagenomic species profiling using universal phylogenetic marker* 1012 *genes.* Nat. Methods, 2013. **10**(12): p. 1196-1199.
- 1013 67. Li, H. and R. Durbin, *Fast and accurate short read alignment with Burrows-Wheeler transform.* 1014 Bioinformatics, 2009. **25**(14): p. 1754-60.
- 101568.Tibshirani, R., Regression Shrinkage and Selection via the Lasso. J.R. Stat. Soc. Series B1016Stat. Methodol., 1996. 58(1): p. 267-288.
- Smialowski, P., D. Frishman, and S. Kramer, *Pitfalls of supervised feature selection*.
 Bioinformatics, 2010. **26**(3): p. 440-3.
- 101970.Benjamini, Y. and Y. Hochberg, Controlling the false discovery rate: a practical and powerful1020approach to multiple testing. J. R. Stat. Soc. Series B Stat. Methodol., 1995. 57(1): p. 289–1021300.
- 102271.Robin, X., et al., pROC: an open-source package for R and S+ to analyze and compare ROC1023curves. BMC Bioinformatics, 2011. **12**(1).
- 1024 72. Oksanen, J., et al., *vegan: Community Ecology Package.* 2018.
- 1025 73. Costea, P.I., et al., *A fair comparison*. Nat. Methods, 2014. **11**(4): p. 359-359.
- 102674.Hastie, T., R. Tibshirani, and J. Friedman, The Elements of Statistical Learning: Data Mining,1027Inference, and Prediction. 2013: Springer Science & Business Media. 536.
- 102875.Peng, H., F. Long, and C. Ding, Feature selection based on mutual information: criteria of1029max-dependency, max-relevance, and min-redundancy. IEEE Trans Pattern Anal Mach Intell,10302005. 27(8): p. 1226-38.

1031 1032	76.	Edgar, R.C., <i>MUSCLE: multiple sequence alignment with high accuracy and high throughput.</i> Nucleic Acids Res., 2004. 32 (5): p. 1792-1797.
1033 1034	77.	Altschul, S.F., et al., <i>Gapped BLAST and PSI-BLAST: a new generation of protein database search programs.</i> Nucleic Acids Res., 1997. 25 (17): p. 3389-3402.
1035 1036	78.	Eddy, S.R., Accelerated Profile HMM Searches. PLoS Comput. Biol., 2011. 7(10): p. e1002195.
1037 1038	79.	Rice, P., I. Longden, and A. Bleasby, <i>EMBOSS: the European Molecular Biology Open Software Suite.</i> Trends Genet., 2000. 16 (6): p. 276-277.
1039 1040 1041	80.	Caporaso, J.G., et al., <i>Global patterns of 16S rRNA diversity at a depth of millions of sequences per sample.</i> Proc Natl Acad Sci U S A, 2011. 108 Suppl 1 : p. 4516-22.





Family





Colorectal cancer metagenomes (Sample positive for this marker species)



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T2D

PD

Other conditions

10% FPR

CTR

0.25

0.00



e

AUROC

8

CD

0

UC

0.5

FR

AT

Abbr.	Condition	Ν
CTR	Healthy controls from meta-analysis	290
T2D	Type 2 diabetes	201
PD	Parkinson 's disease	31
UC	Ulcerative colitis	98
CD	Crohn 's disease	63

CN

US

DE

Classification models based on eggNOG Test set





